

Computing Laboratory

ASSUME-GUARANTEE VERIFICATION FOR PROBABILISTIC SYSTEMS

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Abstract

We present a compositional verification technique for systems that exhibit both probabilistic and nondeterministic behaviour. We adopt an assume-guarantee approach to verification, where both the assumptions made about system components and the guarantees that they provide are regular safety properties, represented by finite automata. Unlike previous proposals for assume-guarantee reasoning about probabilistic systems, our approach does not require that components interact in a fully synchronous fashion. In addition, the compositional verification method is efficient and fully automated, based on a reduction to the problem of multi-objective probabilistic model checking. We present asymmetric and circular assume-guarantee rules, and show how they can be adapted to form quantitative queries, yielding lower and upper bounds on the actual probabilities that a property is satisfied. Our techniques have been implemented and applied to several large case studies, including instances where conventional probabilistic verification is infeasible.

1 Introduction

Many computerised systems exhibit probabilistic behaviour, for example due to the use of randomisation (e.g. in distributed communication or security protocols), or the presence of failures (e.g. in faulty devices or unreliable communication media). The prevalence of such systems in today’s society makes techniques for their formal verification a necessity. This requires models and formalisms that incorporate both *probability* and *nondeterminism*. Although efficient algorithms for verifying such models are known [2, 7] and mature tool support [11, 6] exists, applying these techniques to large, real-life systems remains challenging, and hence techniques to improve scalability are essential.

In this paper, we focus on *compositional* verification techniques for probabilistic and nondeterministic systems, in which a system comprising multiple interacting components can be verified by analysing each component in isolation, rather than verifying the much larger model of the whole system. In the case of *non-probabilistic* models, a successful approach is the use of *assume-guarantee* reasoning. This is based on checking queries of the form $\langle A \rangle M \langle G \rangle$, with the meaning “whenever component M is part of a system satisfying the *assumption* A , then the system is *guaranteed* to satisfy property G ”. Proof rules can then be established that show, for example, that if $\langle true \rangle M_1 \langle A \rangle$ (process M_1 satisfies assumption A in any environment) and $\langle A \rangle M_2 \langle G \rangle$ hold, then the combined system $M_1 \parallel M_2$ satisfies G . For *probabilistic* systems, compositional approaches have also been studied, but a distinct lack of practical progress has been made. In this paper, we address this limitation, presenting the first fully-automated technique for compositional verification of systems exhibiting both probabilistic and nondeterministic behaviour, and illustrating its applicability and efficiency on several large case studies.

We use *probabilistic automata* [19, 20], a well-studied formalism that is naturally suited to modelling multi-component probabilistic systems. Indeed, elegant proof techniques have been developed and used to manually prove correctness of large, complex randomised algorithms [17]. Several branching-time preorders (simulation and bisimulation) have been proposed for probabilistic automata and have been shown to be compositional (i.e.

preserved under parallel composition) [20], but such branching-time equivalences are often too fine to give significant practical advantages for compositional verification.

A coarser linear-time preorder can be obtained through *trace distribution* (probability distributions over sequences of observable actions) inclusion [19]; however, it is well known that this relation is not preserved under parallel composition [18]. Various attempts have been made to characterise refinement relations that *are* preserved, e.g. [19, 14]. An alternative direction is to restrict the forms of parallel composition that are allowed. One example is the formalism of switched probabilistic I/O automata [5], which places restrictions on the scheduling between parallel components. Another is [8] which uses a probabilistic extension of Reactive Modules, restricted to synchronous parallel composition. A limitation of all these approaches is that the relations used, such as trace distribution inclusion and weak probabilistic simulation, are not efficiently computable.

We propose an assume-guarantee verification technique for probabilistic automata, that has no restrictions on the parallel composition permitted between components, allowing greater flexibility to model complex systems. To achieve this, we represent both the assumptions made about system components and the guarantees that they provide as *safety properties*. In the context of probabilistic systems, safety properties capture a wide range of useful properties, e.g. “the maximum probability of an error occurring is at most 0.01” or “the minimum probability of terminating within k time-units is at least 0.75”.

We represent safety properties using finite automata and show that verifying assume-guarantee queries reduces to the problem of *multi-objective model checking* for probabilistic automata [10], which can be implemented efficiently using linear programming. Another key benefit of using finite automata in this way is illustrated by the (non-probabilistic) assume-guarantee verification framework of [15]. There, not only is the verification of queries fully automated, but the assumptions themselves (represented as finite automata) are generated automatically using learning techniques. This opens the way for applying learning techniques to compositional verification in the probabilistic case.

We use our definitions of probabilistic assume guarantee reasoning to formulate and prove several assume-guarantee proof rules, representing commonly occurring patterns of processes. We also discuss how to employ *quantitative* reasoning, in particular obtaining lower and upper bounds on the actual probability that a system satisfies a safety property. The techniques have been implemented in a prototype tool and applied to several large case studies. We demonstrate significant speed-ups over traditional, non-compositional verification, and successfully verify models that cannot be analysed without compositional techniques.

Related work. In addition to the compositional techniques for probabilistic systems surveyed above [5, 8, 14, 17, 18, 19, 20], we mention several other related pieces of work. In particular, our approach was inspired by the large body of work by Giannakopoulou, Pasareanu et al. (see e.g. [15]) on *non-probabilistic* assume guarantee techniques. We also build upon ideas put forward in [10], which suggests using multi-objective verification to check probabilistic assume-guarantee queries. Also relevant are: [9], which presents an

assume/guarantee framework using probabilistic contracts for non-probabilistic models; [3], which presents a theoretical framework for compositional verification of quantitative (but not probabilistic) properties; and [16], which uses probabilistic automata to model the environment of non-probabilistic components.

2 Background

We begin by briefly reviewing probabilistic automata and techniques for their verification. We also introduce safety properties, in the context of probabilistic systems, and discuss multi-objective model checking.

In the following, we use $Dist(S)$ to denote the set of all discrete probability distributions over a set S , η_s for the point distribution on $s \in S$, and $\mu_1 \times \mu_2 \in Dist(S_1 \times S_2)$ for the product distribution of $\mu_1 \in Dist(S_1)$ and $\mu_2 \in Dist(S_2)$.

2.1 Probabilistic automata

Probabilistic automata [19, 20] are a modelling formalism for systems that exhibit both probabilistic and nondeterministic behaviour.

Definition 1 A probabilistic automaton (PA) is a tuple $M = (S, \bar{s}, \alpha_M, \delta_M, L)$ where S is a set of states, $\bar{s} \in S$ is an initial state, α_M is an alphabet, $\delta_M \subseteq S \times (\alpha_M \cup \{\tau\}) \times Dist(S)$ is a probabilistic transition relation and $L : S \rightarrow 2^{AP}$ is a labelling function, assigning atomic propositions from a set AP to each state.

In any state s of a PA M , a *transition*, denoted $s \xrightarrow{a} \mu$, where a is an *action* label and μ is a discrete probability distribution over states, is available¹ if $(s, a, \mu) \in \delta_M$. In an execution of the model, the choice between the available transitions in each state is nondeterministic; the choice of successor state is then made randomly according to the distribution μ . A *path* through M is a (finite or infinite) sequence $s_0 \xrightarrow{a_0, \mu_0} s_1 \xrightarrow{a_1, \mu_1} \dots$ where $s_0 = \bar{s}$ and, for each $i \geq 0$, $s_i \xrightarrow{a_i} \mu_i$ is a transition and $\mu_i(s_{i+1}) > 0$. The sequence of actions a_0, a_1, \dots , after removal of any ‘‘internal actions’’ τ , from a path π is called a *trace* and is denoted $tr(\pi)$.

To reason about PAs, we use the notion of *adversaries* (also called schedulers or strategies), which resolve the nondeterministic choices in a model, based on its execution history. Formally an adversary σ maps any finite path to a sub-distribution over the available transitions in the last state of the path. Adversaries are defined in terms of sub-distributions because they can opt to (with some probability) take none of the available choices and remain in the current state. For this reason, they are sometimes called *partial* adversaries. Occasionally, we will distinguish between these and *complete* adversaries, in which all the distributions are total.

We denote by $Path_M^\sigma$ the set of all paths through M when controlled by adversary σ , and by Adv_M the set of all possible adversaries for M . Under an adversary σ , we

¹Markov decision processes, another commonly used model, are PAs with the restriction that action labels are unique amongst the available transitions for each state.

define a probability space Pr_M^σ over the set of paths $Path_M^\sigma$, which captures the (purely probabilistic) behaviour of M under σ .

To reason about probabilistic systems comprising multiple components, we will need the notions of *parallel composition* and *alphabet extension*:

Definition 2 (Parallel composition of PAs) *If $M_1 = (S_1, \bar{s}_1, \alpha_{M_1}, \delta_{M_1}, L_1)$ and $M_2 = (S_2, \bar{s}_2, \alpha_{M_2}, \delta_{M_2}, L_2)$ are PAs, then their parallel composition, denoted $M_1 \parallel M_2$, is given by the PA $(S_1 \times S_2, (\bar{s}_1, \bar{s}_2), \alpha_{M_1} \cup \alpha_{M_2}, \delta_{M_1 \parallel M_2}, L)$ where $\delta_{M_1 \parallel M_2}$ is defined such that $(s_1, s_2) \xrightarrow{a} \mu_1 \times \mu_2$ if and only if one of the following holds:*

- $s_1 \xrightarrow{a} \mu_1, s_2 \xrightarrow{a} \mu_2$ and $a \in \alpha_{M_1} \cap \alpha_{M_2}$
- $s_1 \xrightarrow{a} \mu_1, \mu_2 = \eta_{s_2}$ and $a \in (\alpha_{M_1} \setminus \alpha_{M_2}) \cup \{\tau\}$
- $s_2 \xrightarrow{a} \mu_2, \mu_1 = \eta_{s_1}$ and $a \in (\alpha_{M_2} \setminus \alpha_{M_1}) \cup \{\tau\}$

and $L(s_1, s_2) = L_1(s_1) \cup L_2(s_2)$.

Definition 3 (Alphabet extension) *For any PA $M = (S, \bar{s}, \alpha_M, \delta_M, L)$ and set of actions Σ , we extend the alphabet of M to Σ , denoted $M[\Sigma]$, as follows: $M[\Sigma] = (S, \bar{s}, \alpha_M \cup \Sigma, \delta_{M[\Sigma]}, L)$ where $\delta_{M[\Sigma]} = \delta_M \cup \{(s, a, \eta_s) \mid s \in S \wedge a \in \Sigma \setminus \alpha_M\}$.*

We also require the notion of *projections*. First, for any state $s = (s_1, s_2)$ of $M_1 \parallel M_2$, the projection of s onto M_i , denoted by $s \upharpoonright_{M_i}$, is s_i . We extend this notation to distributions over the state space $S_1 \times S_2$ of $M_1 \parallel M_2$ in the standard manner. Next, for any path π of $M_1 \parallel M_2$, the projection of π onto M_i , denoted $\pi \upharpoonright_{M_i}$, is the path obtained from π by projecting each state of π onto M_i and removing all the actions not in α_{M_i} together with the subsequent states.

Definition 4 (Projections of adversaries) *Let M_1 and M_2 be PAs and σ an adversary of $M_1 \parallel M_2$. The projection of σ onto M_i , denoted $\sigma \upharpoonright_{M_i}$, is the adversary on M_i where, for any finite path π of M_i :*

$$\sigma \upharpoonright_{M_i}(\pi)(a, \mu) = \sum \{Pr(\pi') \cdot \sigma(\pi')(a, \mu') \mid \pi' \in Path_{M_1 \parallel M_2}^\sigma \wedge \pi' \upharpoonright_{M_i} = \pi \wedge \mu' \upharpoonright_{M_i} = \mu\}.$$

Compositional reasoning about PAs, and in particular adversary projections, necessitates the use of partial, rather than complete, adversaries. In particular, even if an adversary σ of $M_1 \parallel M_2$ is complete, the projection $\sigma \upharpoonright_{M_i}$ onto one component may be partial.

2.2 Model checking for PAs

The verification of PAs against properties specified either in temporal logic or as automata has been well studied. In this paper, both the states and transitions of PAs are labelled (with sets of atomic propositions and actions, respectively) and we formulate properties that refer to both types of labels. For the former, we will express properties in linear temporal logic (LTL), and for the latter, we will use safety properties represented by deterministic finite automata.

LTL verification. For an LTL formula ψ , PA M and adversary $\sigma \in Adv_M$:

$$Pr_M^\sigma(\psi) \stackrel{\text{def}}{=} Pr_M^\sigma\{\pi \in Path_M^\sigma \mid \pi \models \psi\}$$

where $\pi \models \psi$ denotes satisfaction according to the standard semantics of LTL. Verifying an LTL specification ψ against M typically involves checking that the probability of satisfying ψ meets a probability bound for all adversaries. This reduces to computing the minimum or maximum probability of satisfying ψ :

$$Pr_M^{\min}(\psi) \stackrel{\text{def}}{=} \inf_{\sigma \in Adv_M} Pr_M^\sigma(\psi) \quad \text{and} \quad Pr_M^{\max}(\psi) \stackrel{\text{def}}{=} \sup_{\sigma \in Adv_M} Pr_M^\sigma(\psi).$$

The complexity of this computation is polynomial in the size of M and doubly exponential in the size of ψ [7]. In practice, the LTL formula ψ is small and, for simple, commonly used cases such as $\diamond ap$ (“eventually ap ”) or $\square ap$ (“globally ap ”), model checking is polynomial [2]. Furthermore, efficient implementations of LTL verification exist in tools such as PRISM [11] and LiQuor [6].

Safety properties. A *regular safety property* A represents a set of infinite words, denoted $\mathcal{L}(A)$, that is characterised by a regular language of *bad prefixes*, finite words of which any extension is not in $\mathcal{L}(A)$. More precisely, we will define a regular safety property A by a (complete) deterministic finite automaton (DFA) $A^{err} = (Q, \bar{q}, \alpha_A, \delta_A, F)$, comprising states Q , initial state $\bar{q} \in Q$, alphabet α_A , transition function $\delta_A : Q \times \alpha_A \rightarrow Q$ and accepting states $F \subseteq Q$. The DFA A^{err} defines, in standard fashion, a regular language $\mathcal{L}(A^{err}) \subseteq (\alpha_A)^*$. The language $\mathcal{L}(A)$ is then defined as $\mathcal{L}(A) = \{w \in (\alpha_A)^\omega \mid \text{no prefix of } w \text{ is in } \mathcal{L}(A^{err})\}$.

Given a PA M , adversary $\sigma \in Adv_M$ and regular safety property A with $\alpha_A \subseteq \alpha_M$, we define the probability of M under σ satisfying A as:

$$Pr_M^\sigma(A) \stackrel{\text{def}}{=} Pr_M^\sigma\{\pi \in Path_M^\sigma \mid tr(\pi) \upharpoonright_{\alpha_A} \in \mathcal{L}(A)\}$$

where $w \upharpoonright_\alpha$ is the projection of word w onto a subset α of its alphabet. We then define $Pr_M^{\min}(A)$ and $Pr_M^{\max}(A)$ as for LTL above.

Definition 5 (Probabilistic safety properties) A probabilistic safety property $\langle A \rangle_{\geq p}$ comprises a regular safety property A and a rational probability bound p . We say that a PA M satisfies the property, denoted $M \models \langle A \rangle_{\geq p}$, if the probability of satisfying A is at least p for any adversary:

$$M \models \langle A \rangle_{\geq p} \Leftrightarrow \forall \sigma \in Adv_M . Pr_M^\sigma(A) \geq p \Leftrightarrow Pr_M^{\min}(A) \geq p.$$

Safety properties can be used to represent a wide range of useful properties of probabilistic automata. Examples include:

- “the probability of an error occurring is at most 0.01”
- “event A always occurs before event B with probability at least 0.98”
- “the probability of terminating within k time-units is at least 0.75”

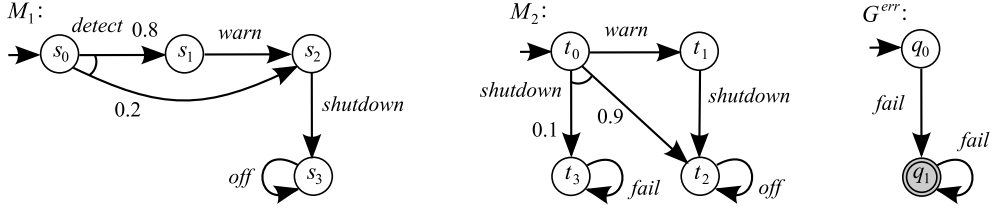


Figure 1: Two probabilistic automata M_1, M_2 and the DFA for a safety property G

The last of these represents a very useful class of properties for *timed* probabilistic systems, perhaps not typically considered as safety properties. Using the *digital clocks* approach of [12], verifying real-time probabilistic systems can often be reduced to analysis of a PA with time steps encoded as a special action type. Such requirements are then naturally encoded as safety properties.

Example 1. Figure 1 shows two PAs M_1 and M_2 . Component M_1 represents a controller that powers down devices. Upon receipt of the *detect* signal, it first issues the *warn* signal followed by *shutdown*; however, with probability 0.2 it will fail to issue the *warn* signal. M_2 represents a device which, given the *shutdown* signal, powers down correctly if it first receives the *warn* signal and otherwise will only power down correctly 90% of the time. We consider a simple safety property G “action *fail* never occurs”, represented by the DFA G^{err} also shown in Figure 1. Composing the two PAs in parallel and applying model checking, we have that $Pr_{M_1 \parallel M_2}^{\min}(G) = 0.98$. Thus, $M_1 \parallel M_2 \models \langle G \rangle_{\geq 0.98}$.

Safety verification. Using standard automata-based techniques for model checking PAs [7], verifying correctness of probabilistic safety properties reduces to model checking the product of a PA and a DFA:

Definition 6 (PA-DFA product) *The product of a PA $M = (S, \bar{s}, \alpha_M, \delta_M, L)$ and DFA $A^{err} = (Q, \bar{q}, \alpha_A, \delta_A, F)$ with $\alpha_A \subseteq \alpha_M$ is given by the PA $M \otimes A^{err} = (S \times Q, (\bar{s}, \bar{q}), \alpha_M, \delta', L')$ where:*

- $(s, q) \xrightarrow{a} \mu \times \eta_{q'}$ if $s \xrightarrow{a} \mu$ and $q' = \delta_A(q, a)$ if $a \in \alpha_A$ and $q' = q$ otherwise;
- $L'(s, q) = L(s) \cup \{err_A\}$ if $q \in F$ and $L'(s, q) = L(s)$ otherwise.

Proposition 2.1 *For PA M and regular safety property A , we have:*

$$Pr_M^{\min}(A) = 1 - Pr_{M \otimes A^{err}}^{\max}(\diamond err_A).$$

Thus, using [2], satisfaction of the probabilistic safety property $\langle A \rangle_{\geq p}$ can be checked in time polynomial in the size of $M \otimes A^{err}$. Note that maximum reachability probabilities, and therefore satisfaction of probabilistic safety properties, are independent of whether complete or partial adversaries are considered.

Multi-objective model checking. In addition to traditional probabilistic model checking techniques, the approach presented in this paper requires the use of *multi-objective* model checking [10]. The conventional approach described above allows us to check whether, for all adversaries (or, dually, for at least one adversary), the probability of some property is above (or below) a given bound. Multi-objective queries allow us to check the existence of an adversary satisfying *multiple* properties of this form. In particular, consider k predicates of the form $Pr_M^\sigma(\psi_i) \sim_i p_i$ where ψ_i is an LTL formula, $p_i \in [0, 1]$ is a rational probability bound and $\sim_i \in \{\geq, >\}$. Using the techniques in [10], we can verify whether:

$$\exists \sigma \in Adv_M . \bigwedge_{i=1}^k (Pr_M^\sigma(\psi_i) \sim_i p_i)$$

by a reduction to a linear programming (LP) problem. Like for (single-objective) LTL verification, this can be done in time polynomial in the size of M (and doubly exponential in the sizes of ψ_i). In fact, [10] also shows that this technique generalises to checking existential or universal queries over a Boolean combination of predicates for which $\sim_i \in \{\geq, >, \leq, <\}$. In all cases, if an adversary which satisfies the predicates exists, then it can also easily be obtained.

Finally, through a trivial extension of this approach (and without increasing the complexity), we can formulate *quantitative* multi-objective queries. For example, given a conjunction of the above predicates $\Psi = \bigwedge_{i=1}^k Pr_M^\sigma(\psi_i) \sim_i p_i$, and an additional LTL formula ψ_0 , we can compute the maximum probability of ψ_0 that is achievable whilst also satisfying Ψ :

$$Pr_M^{\max}(\psi_0 | \Psi) \stackrel{\text{def}}{=} \sup\{Pr_M^\sigma(\psi_0) | \sigma \in Adv_M \wedge \Psi\}.$$

3 Compositional Verification for PAs

We now describe our approach for compositional verification of probabilistic automata. We first define the basic underlying ideas and then present several different proof rules. For clarity, we present the simplest of these rules in some detail and then discuss some generalisations and extensions.

We extend the notion of *assume-guarantee* reasoning to PAs using *probabilistic assume-guarantee triples* of the form $\langle A \rangle_{\geq p_A} M \langle G \rangle_{\geq p_G}$, where $\langle A \rangle_{\geq p_A}$ and $\langle G \rangle_{\geq p_G}$ are probabilistic safety properties and M is a PA. Informally, the meaning of this is “whenever M is part of a system satisfying A with probability at least p_A , then the system will satisfy G with probability at least p_G ”. Formally:

Definition 7 (Assume-guarantee semantics) *If $\langle A \rangle_{\geq p_A}$ and $\langle G \rangle_{\geq p_G}$ are probabilistic safety properties, M is a PA and $\alpha_G \subseteq \alpha_A \cup \alpha_M$, then*

$$\langle A \rangle_{\geq p_A} M \langle G \rangle_{\geq p_G} \Leftrightarrow \forall \sigma \in Adv_{M[\alpha_A]} . \left(Pr_{M[\alpha_A]}^\sigma(A)_{\geq p_A} \rightarrow Pr_{M[\alpha_A]}^\sigma(G)_{\geq p_G} \right).$$

The use of $M[\alpha_A]$, i.e. M extended to the alphabet of A , in this definition is required for the case where the property G includes actions that are not in M .

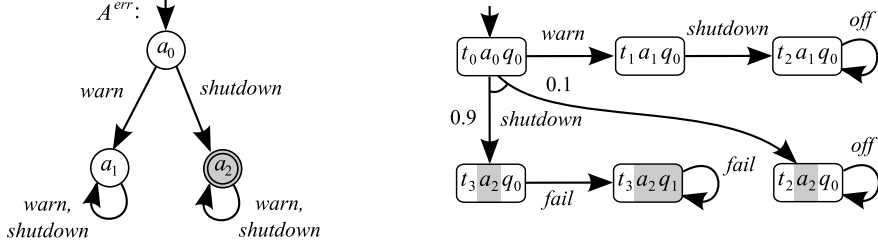


Figure 2: DFA for safety property A and the product PA $M_2 \otimes A^{err} \otimes G^{err}$ (see Figure 1)

We write $\langle true \rangle M \langle G \rangle_{\geq p_G}$ to denote the absence of any assumption, i.e. the query $\langle true \rangle M \langle G \rangle_{\geq p_G}$ is equivalent to $M \models \langle G \rangle_{\geq p_G}$ which, as described above, is standard model checking [2]. In the general case, we check the satisfaction of a probabilistic assume-guarantee triple using multi-objective PA model checking:

Proposition 3.1 (Assume-guarantee model checking) *Let M be a PA, $\langle A \rangle_{\geq p_A}$, $\langle G \rangle_{\geq p_G}$ be probabilistic safety properties and $M' = M[\alpha_A] \otimes A^{err} \otimes G^{err}$. The probabilistic assume-guarantee triple $\langle A \rangle_{\geq p_A} M \langle G \rangle_{\geq p_G}$ holds if and only if:*

$$\neg \exists \sigma' \in Adv_{M'} . \left(Pr_{M'}^{\sigma'}(\Box \neg err_A) \geq p_A \wedge Pr_{M'}^{\sigma'}(\Diamond err_G) > 1 - p_G \right)$$

which can be checked in time polynomial in $|M'|$ by solving an LP problem [10].

We now present, using the definitions above, several assume-guarantee proof rules to allow compositional verification.

An asymmetric proof rule. The first rule we consider is *asymmetric*, in the sense that we require only a single assumption about one component. Experience in the non-probabilistic setting [15] indicates that, despite its simplicity, rules of this form are widely applicable.

Theorem 3.2 *If M_1, M_2 are probabilistic automata and $\langle A \rangle_{\geq p_A}, \langle G \rangle_{\geq p_G}$ probabilistic safety properties such that $\alpha_A \subseteq \alpha_{M_1}$ and $\alpha_G \subseteq \alpha_{M_2} \cup \alpha_A$, then the following proof rule holds:*

$$\frac{\langle true \rangle M_1 \langle A \rangle_{\geq p_A} \quad \langle A \rangle_{\geq p_A} M_2 \langle G \rangle_{\geq p_G}}{\langle true \rangle M_1 \parallel M_2 \langle G \rangle_{\geq p_G}} \quad (\text{ASYM})$$

Theorem 3.2 means that, given an appropriate assumption $\langle A \rangle_{\geq p_A}$, we can check the correctness of a probabilistic safety property $\langle G \rangle_{\geq p_G}$ on $M_1 \parallel M_2$, without constructing and model checking the full model. Instead, we perform one instance of (standard) model checking on M_1 (to check the first condition of rule (ASYM)) and one instance of multi-objective model checking on $M_2[\alpha_A] \otimes A^{err}$ (to check the second). If A^{err} is much smaller than M_1 , we can expect significant gains in terms of the verification performance.

Example 2. We illustrate the rule (ASYM) on the PAs M_1, M_2 and property $\langle G \rangle_{\geq 0.98}$ from Example 1. Figure 2 (left) shows a DFA A^{err} representing the safety property A

“*warn* occurs before *shutdown*”. We will use the probabilistic safety property $\langle A \rangle_{\geq 0.8}$ as the assumption about M_1 in (ASYM).

Checking the first condition of (ASYM) amounts to verifying $M_1 \models \langle A \rangle_{\geq 0.8}$, which can be done with standard probabilistic model checking. To complete the verification, we need to check the second condition $\langle A \rangle_{\geq 0.8} M_2 \langle G \rangle_{\geq 0.98}$, which, from Proposition 3.1, is achieved though multi-objective model checking on the product² $M_2 \otimes A^{err} \otimes G^{err}$. More precisely, we check there is no adversary under which the probability of remaining within states not satisfying err_A is at least 0.8 and the probability of reaching an err_G state is above $1 - 0.98 = 0.02$. The product is shown in Figure 2 (right), where we indicate states satisfying err_A and err_G by highlighting the accepting states a_2 and q_1 of DFAs A^{err} and G^{err} . By inspection, we see that no such adversary exists, so we can conclude that $M_1 \parallel M_2 \models \langle G \rangle_{\geq 0.98}$. Consider, however, the adversary σ which, in the initial state, chooses *warn* with probability 0.8 and *shutdown* with probability 0.2. This satisfies $\Box \neg err_A$ with probability 0.8 and $\Diamond err_G$ with probability 0.02. Hence, $\langle A \rangle_{\geq 0.8} M_2 \langle G \rangle_{\geq p_G}$ does *not* hold for any value of $p_G > 1 - 0.02 = 0.98$.

Proof of Theorem 3.2. We give below the proof of Theorem 3.2. This requires the following lemma, which is a simple extension of [19, Lemma 7.2.6, page 141].

Lemma 3.3 *Let M_1, M_2 be PAs, $\sigma \in Adv_{M_1 \parallel M_2}$, $\Sigma \subseteq \alpha_{M_1 \parallel M_2}$ and $i = 1, 2$. If A and B are regular safety properties such that $\alpha_A \subseteq \alpha_{M_i}$ and $\alpha_B \subseteq \alpha_{M_i[\Sigma]}$, then*

$$(a) Pr_{M_1 \parallel M_2}^\sigma(A) = Pr_{M_i}^{\sigma \upharpoonright_{M_i}}(A) \quad \text{and} \quad (b) Pr_{M_1 \parallel M_2}^\sigma(B) = Pr_{M_i[\Sigma]}^{\sigma \upharpoonright_{M_i[\Sigma]}}(B).$$

Note that the projections onto $M_i[\Sigma]$ in the above are well defined since the condition $\Sigma \subseteq \alpha_{M_1 \parallel M_2}$ implies that $M_1 \parallel M_2 = M_1[\Sigma] \parallel M_2 = M_1 \parallel M_2[\Sigma]$.

Proof 3.4 (of Theorem 3.2) *The proof is by contradiction. Assume that there exist PAs M_1 and M_2 and probabilistic safety properties $\langle A \rangle_{\geq p_A}$ and $\langle G \rangle_{\geq p_G}$ such that $\langle true \rangle M_1 \langle A \rangle_{\geq p_A}$ and $\langle A \rangle_{\geq p_A} M_2 \langle G \rangle_{\geq p_G}$ hold, while $\langle true \rangle M_1 \parallel M_2 \langle G \rangle_{\geq p_G}$ does not. From the latter, it follows that there exists an adversary $\sigma \in Adv_{M_1 \parallel M_2}$ such that $Pr_{M_1 \parallel M_2}^\sigma(G) < p_G$. Now, since $\langle true \rangle M_1 \langle A \rangle_{\geq p_A}$ and $\sigma \upharpoonright_{M_1} \in Adv_{M_1}$, it follows that:*

$$\begin{aligned} Pr_{M_1}^{\sigma \upharpoonright_{M_1}}(A) \geq p_A &\Rightarrow Pr_{M_1 \parallel M_2}^\sigma(A) \geq p_A && \text{by Lemma 3.3(a) since } \alpha_A \subseteq \alpha_{M_1} \\ &\Rightarrow Pr_{M_2[\alpha_A]}^{\sigma \upharpoonright_{M_2[\alpha_A]}}(A) \geq p_A && \text{by Lemma 3.3(b) since } \alpha_A \subseteq \alpha_{M_2[\alpha_A]} \\ &\Rightarrow Pr_{M_2[\alpha_A]}^{\sigma \upharpoonright_{M_2[\alpha_A]}}(G) \geq p_G && \text{since } \langle A \rangle_{\geq p_A} M_2 \langle G \rangle_{\geq p_G} \\ &\Rightarrow Pr_{M_1 \parallel M_2}^\sigma(G) \geq p_G && \text{by Lemma 3.3(b) since } \alpha_G \subseteq \alpha_{M_2[\alpha_A]} \end{aligned}$$

which contradicts the assumption that $Pr_{M_1 \parallel M_2}^\sigma(G) < p_G$. □

²In this example, $\alpha_A = \{\text{warn}, \text{shutdown}\} \subseteq \alpha_{M_2}$ so $M_2[\alpha_A] = M_2$.

Generalising the proof rule. Next, we state two useful generalisations of the above proof rule. First, using $\langle A_1, \dots, A_k \rangle_{\geq p_1, \dots, p_k}$ to denote the conjunction of probabilistic safety properties $\langle A_i \rangle_{\geq p_i}$ for $i = 1, \dots, k$, we have:

$$\frac{\langle true \rangle M_1 \langle A_1, \dots, A_k \rangle_{\geq p_1, \dots, p_k} \quad \langle A_1, \dots, A_k \rangle_{\geq p_1, \dots, p_k} M_2 \langle G \rangle_{\geq p_G}}{\langle true \rangle M_1 \parallel M_2 \langle G \rangle_{\geq p_G}} \quad (\text{ASYM-MULT})$$

Definition 7 extends naturally to k assumptions, replacing α_A with $\cup_{i=1}^k \alpha_{A_i}$ and the single probabilistic safety property on the left-hand side of the implication with the conjunction. In similar fashion, by adapting Proposition 3.1, model checking of the query $\langle A_1, \dots, A_k \rangle_{\geq p_1, \dots, p_k} M \langle G \rangle_{\geq p_G}$ reduces to multi-objective model checking on the product $M[\cup_{i=1}^k \alpha_{A_i}] \otimes A_1^{err} \otimes \dots \otimes A_k^{err} \otimes G^{err}$.

Secondly, we observe that, through repeated application of (ASYM), we obtain a rule of the following form for n components:

$$\frac{\langle true \rangle M_1 \langle A_1 \rangle_{\geq p_1} \quad \langle A_1 \rangle_{\geq p_1} M_2 \langle A_2 \rangle_{\geq p_2} \quad \dots \quad \langle A_{n-1} \rangle_{\geq p_{n-1}} M_n \langle G \rangle_{\geq p_G}}{\langle true \rangle M_1 \parallel \dots \parallel M_n \langle G \rangle_{\geq p_G}} \quad (\text{ASYM-N})$$

A circular proof rule. One potential limitation of the rule (ASYM) is that we may not be able to show that the assumption A_1 about M_1 holds without making additional assumptions about M_2 . This can be overcome by using the following *circular* proof rule:

Theorem 3.5 *If M_1, M_2 are PAs and $\langle A_1 \rangle_{\geq p_1}, \langle A_2 \rangle_{\geq p_2}$ and $\langle G \rangle_{\geq p_G}$ probabilistic safety properties such that $\alpha_{A_2} \subseteq \alpha_{M_2}, \alpha_{A_1} \subseteq \alpha_{M_1} \cup \alpha_{A_2}$ and $\alpha_G \subseteq \alpha_{M_2} \cup \alpha_{A_1}$, then the following circular assume-guarantee proof rule holds:*

$$\frac{\langle true \rangle M_2 \langle A_2 \rangle_{\geq p_2} \quad \langle A_2 \rangle_{\geq p_2} M_1 \langle A_1 \rangle_{\geq p_1} \quad \langle A_1 \rangle_{\geq p_1} M_2 \langle G \rangle_{\geq p_G}}{\langle true \rangle M_1 \parallel M_2 \langle G \rangle_{\geq p_G}} \quad (\text{CIRC})$$

An asynchronous proof rule. This rule is motivated by the fact that, often, part of a system comprises several *asynchronous* components, that is, components with disjoint alphabets. In such cases, it can be difficult to establish useful probability bounds on the combined system if the fact that the components act *independently* is ignored. For example, consider the case of n independent coin flips; in isolation, we have that the probability of any coin not returning a tail is $1/2$. Now, ignoring the independence of the coins, all we can say is that the probability of any of them not returning a tail is at least $1/2$. However, using their independence, we have that this probability is at least $1 - 1/2^n$.

Theorem 3.6 For any PAs M_1, M_2 and probabilistic safety properties $\langle A_1 \rangle_{\geq p_{A_2}}, \langle A_2 \rangle_{\geq p_{A_1}}, \langle G_1 \rangle_{\geq p_{G_1}}$ and $\langle G_2 \rangle_{\geq p_{G_2}}$ such that $\alpha_{M_1} \cap \alpha_{M_2} = \emptyset$, $\alpha_{G_1} \subseteq \alpha_{M_1} \cup \alpha_{A_1}$ and $\alpha_{G_2} \subseteq \alpha_{M_2} \cup \alpha_{A_2}$, we have the following asynchronous assume-guarantee proof rule:

$$\frac{\begin{array}{c} \langle A_1 \rangle_{\geq p_{A_1}} M_1 \langle G_1 \rangle_{\geq p_{G_1}} \\ \langle A_2 \rangle_{\geq p_{A_2}} M_2 \langle G_2 \rangle_{\geq p_{G_2}} \end{array}}{\langle A_1, A_2 \rangle_{\geq p_{A_1}, p_{A_2}} M_1 \| M_2 \langle G_1 \vee G_2 \rangle_{\geq p_{G_1} + p_{G_2} - p_{G_1} \cdot p_{G_2}}} \quad (\text{ASYNC})$$

where the disjunction of safety properties G_1 and G_2 is obtained by taking the intersection of the DFAs G_1^{err} and G_2^{err} .

4 Quantitative Assume-Guarantee Queries

Practical experience with probabilistic verification suggests that it is often more useful to adopt a *quantitative* approach. For example, rather than checking the correctness of a probabilistic safety property $\langle G \rangle_{\geq p_G}$, it may be preferable to just compute the actual worst-case (minimum) probability $Pr_M^{\min}(G)$ that G is satisfied. In this section we consider how to formulate such quantitative queries in the context of assume-guarantee reasoning. For simplicity, we restrict our attention here to the rule (ASYM) for fixed PAs M_1 and M_2 , and property G . Similar reasoning applies to the other rules presented above.

Maximal lower bounds. Rule (ASYM) allows us to establish *lower bounds* for the probability $Pr_{M_1 \| M_2}^{\min}(G)$, i.e. it can be used to prove, for certain values of p_G , that $Pr_{M_1 \| M_2}^{\min}(G) \geq p_G$. We consider now how to obtain the highest such lower bound, say p_G^* . First, we note that, from Definition 7, it is clear that the highest value of p_G for which $\langle A \rangle_{\geq p_A} M_2 \langle G \rangle_{\geq p_G}$ holds will be obtained by using the maximum possible value of p_A . For rule (ASYM) to be applicable, this is equal to $Pr_{M_1}^{\min}(A)$, since for any higher value of p_A the first condition will fail to hold. Now, by Proposition 3.1, and letting $M' = M_2[\alpha_A] \otimes A^{err} \otimes G^{err}$, the value p_G^* can be obtained through multi-objective model checking as follows:

$$p_G^* = 1 - Pr_{M'}^{\max}(\diamond err_G \mid \Psi) \text{ where } \Psi = Pr_{M'}^{\sigma}(\square \neg err_A) \geq p_A.$$

Parameterised queries. Let us assume that component M_1 is parameterised by a variable x in such a way that varying x changes the probability of M_1 satisfying the assumption A . For example, increasing the value of x might increase the probability $Pr_{M_1}(A)$, but simultaneously worsen some other performance measure or cost associated with M_1 . In this situation, it is desirable to establish a trade-off between the probability of $M_1 \| M_2$ satisfying G and the secondary ‘cost’ of M_1 . Our use of multi-objective model checking for compositional verification offers two choices here. Firstly, we can pick a suitable threshold for $Pr_{M_1 \| M_2}(G)$ and then compute the lowest value of $Pr_{M_1}(A)$ which guarantees this, allowing an appropriate value of x to be chosen. Alternatively, we can consider the so-called *Pareto curve*: the set of achievable combinations of $Pr_{M_1 \| M_2}(G)$

and $Pr_{M_1}(A)$, which will present a clear view of the trade-off. For the latter, we can use the techniques of [10] for approximate exploration of the Pareto curve.

Upper bounds. Since application of (ASYM) gives lower bounds on $Pr_{M_1||M_2}^{\min}(G)$, it is desirable to also generate *upper* bounds on this probability. This can be done as follows. When checking condition 2 of (ASYM), using multi-objective model checking, we also obtain an adversary $\sigma \in Adv_{M_2[\alpha_A] \otimes A^{err}}$ that satisfies $\langle A \rangle_{\geq p_A}$ and gives the minimum (i.e. worst-case) probability of satisfying G . This can then be projected onto M_2 , giving an adversary σ_2 which achieves the worst-case behaviour of the single component M_2 with respect to G satisfying $\langle A \rangle_{\geq p_A}$. Furthermore, from σ_2 , we can easily construct a PA $M_2^{\sigma_2}$ that represents the behaviour of M_2 under σ_2 .

Finally, we compute the probability of satisfying G on $M_1||M_2^{\sigma_2}$. Because $M_2^{\sigma_2}$ is likely to be much smaller than M_2 , there is scope for this to be efficient, even if model checking $M_1||M_2$ in full is not feasible. Since $M_1||M_2^{\sigma_2}$ represents only a subset of the behaviour of $M_1||M_2$, the probability computed is guaranteed to give an upper bound on $Pr_{M_1||M_2}^{\min}(G)$. We use σ_2 (which achieves the worst-case behaviour with respect to G), rather than an arbitrary adversary of M_2 , in order to obtain a tighter upper bound.

5 Implementation and Case Studies

We have implemented our compositional verification approach in a prototype tool. Recall that, using the rules given in Section 3, verification requires both standard (automata-based) model checking and multi-objective model checking. Our tool is based on the probabilistic model checker PRISM [11], which already supports LTL model checking of probabilistic automata. Model checking of probabilistic safety properties, represented by DFAs, can be achieved with existing versions of PRISM, since DFAs can easily be encoded in PRISM’s modelling language. For multi-objective model checking, we have extended PRISM with an implementation of the techniques in [10]. This requires the solution of Linear Programming (LP) problems, for which we use the ECLiPSe Constraint Logic Programming system with the COIN-OR CBC solver, implementing a branch-and-cut algorithm. All experiments were run on a 2GHz PC with 2GB RAM. Any run exceeding a time-limit of 24 hours was disregarded.

We demonstrate the application of our tool to two large case studies. The first is the randomised consensus algorithm of Aspnes & Herlihy [1]. The algorithm allows N processes in a distributed network to reach a consensus and employs, in each round, a shared coin protocol parameterised by K . The PA model is based on [13] and consists of an automaton for each process and for the shared coin protocol of each round. We analyse the minimum probability that the processes decide by round R . The compositional verification employs $R-2$ uses of the ASYNC rule to return a probabilistic safety property satisfied by the (asynchronous) composition of the shared coin protocols for the first $R-2$ rounds. This is then used as the assumption of an ASYM rule for the subsystem representing the processes.

The second case study is the Zeroconf network configuration protocol [4]. We use the

Case study [parameters]		Non-compositional			Compositional		
		States	Time (s)	Result [†]	LP size	Time (s)	Result [†]
<i>consensus</i> (2 <i>processes</i>) [<i>R K</i>]	3 2	5,158	1.6	0.108333	1,064	0.9	0.108333
	3 20	40,294	108.1	0.012500	1,064	7.4	0.012500
	4 2	20,886	3.6	0.011736	2,372	1.2	0.011736
	4 20	166,614	343.1	0.000156	2,372	7.8	0.000156
	5 2	83,798	7.7	0.001271	4,988	2.2	0.001271
	5 20	671,894	1,347	0.000002	4,988	8.8	0.000002
<i>consensus</i> (3 <i>processes</i>) [<i>R K</i>]	3 2	1,418,545	18,971	0.229092	40,542	29.6	0.229092
	3 12	16,674,145*	time-out	-	40,542	49.7	0.041643
	3 20	39,827,233*	time-out	-	40,542	125.3	0.024960
	4 2	150,487,585	78,955	0.052483	141,168	376.1	0.052483
	4 12	1,053,762,385*	mem-out	-	141,168	396.3	0.001734
	4 20	2,028,200,209*	mem-out	-	141,168	471.9	0.000623
<i>zeroconf</i> [<i>K</i>]	2	91,041	39.0	2.0e-5	6,910	9.3	3.1e-4
	4	313,541	103.9	7.3e-7	20,927	21.9	3.1e-4
	6	811,290	275.2	2.6e-8	40,258	54.8	2.5e-4
	8	1,892,952	592.2	9.5e-10	66,436	107.6	9.0e-6
<i>zeroconf</i> (<i>time</i> <i>bounded</i>) [<i>K T</i>]	2 10	665,567	46.3	5.9e-5	62,188	89.0	2.1e-4
	2 14	106,177	63.1	2.0e-8	101,313	170.8	8.1e-8
	4 10	976,247	88.2	3.3e+0	74,484	170.8	3.3e+0
	4 14	2,288,771	128.3	7.0e-5	166,203	430.6	3.1e-4

* These models can be constructed, but not model checked, in PRISM.

[†] Results are maximum probabilities of error so actual values are these subtracted from 1.

Table 1: Experimental results, comparing with non-compositional verification

PA model from [12] consisting of two components, one representing a new host joining the network (parameterised by K , the number of probes it sends before using an IP address), and the second representing the environment, i.e. the existing network. We consider two properties: the minimum probability that a host employs a fresh IP address and that a host is configured by time T . In each case the compositional verification uses one application of the CIRC rule.

Table 1 shows experimental results for these case studies. We present the total time required for both compositional verification, as described in this paper, and non-compositional verification using PRISM (with the fastest available engine). Note that, in each case, we use the quantitative approach described in Section 4 and give actual (bounds on) probabilities computed. To give an indication of the size of the models considered, we give the number of states for the full (non-compositional) models and the number of variables in the LP problems used for multi-objective model checking in the compositional case.

In summary, we see that the compositional approach is faster in the majority of cases. Furthermore, it allows verification of several models for which it is infeasible with conventional techniques. For the cases where compositional verification is slower, this is

due to the cost of solving a large LP problem, which is known to be more expensive than the highly optimised techniques used in PRISM. Furthermore, LP solution represents the limiting factor with respect to the scalability of the compositional approach. We expect that improvements to our technique can be made that will reduce LP problem sizes and improve performance. Finally, we note that the numerical values produced using compositional verification are generally good; in fact, for the consensus case study, the bounds obtained are precise.

6 Conclusions

We have presented a compositional verification technique, based on assume-guarantee rules, for probabilistic automata. Properties of these models are represented as probabilistic safety properties, and we show how verifying the resulting assume-guarantee queries reduces to the problem of multi-objective model checking. We also show how this can be leveraged to provide a *quantitative* approach to compositional verification. In contrast to existing work in this area, our techniques can be implemented efficiently and we demonstrate successful results on several large case studies.

There are several interesting directions for future work. In particular, we plan to experiment with the use of learning techniques to automatically produce the assumptions required for compositional reasoning. We also intend to further develop our compositional proof rules and investigate to what extent they are complete. Finally, we plan to expand the range of properties that can be verified, including for example reward-based specifications.

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Appendix.

We include here proofs omitted from the main text.

Proof .1 (of Proposition 3.1) Consider any PA M , probabilistic safety properties $\langle A \rangle_{\geq p_A}$ and $\langle G \rangle_{\geq p_G}$, and let $M' = Adv_{M[\alpha_A]} \otimes A^{err} \otimes G^{err}$. By definition of \otimes (see Definition 6) we can construct a bijective function $f : Adv_{M[\alpha_A]} \rightarrow Adv_{M'}$ such that:

$$Pr_{M[\alpha_A]}^\sigma(B) = 1 - Pr_{M'}^{f(\sigma)}(\diamond err_B) \text{ for } B \in \{A, G\}. \quad (1)$$

Now using Definition 7 we have:

$$\begin{aligned} \langle A \rangle_{\geq p_A} M \langle G \rangle_{\geq p_G} &\Leftrightarrow \forall \sigma \in Adv_{M[\alpha_A]}. \left(Pr_{M[\alpha_A]}^\sigma(A)_{\geq p_A} \rightarrow Pr_{M[\alpha_A]}^\sigma(G)_{\geq p_G} \right) \\ &\Leftrightarrow \forall \sigma \in Adv_{M[\alpha_A]}. \left(Pr_{M'}^{f(\sigma)}(\diamond err_A) \leq 1 - p_A \rightarrow Pr_{M'}^{f(\sigma)}(\diamond err_G) \leq 1 - p_G \right) \\ &\Leftrightarrow \forall \sigma' \in Adv_{M'}. \left(Pr_{M'}^{\sigma'}(\diamond err_A) \leq 1 - p_A \rightarrow Pr_{M'}^{\sigma'}(\diamond err_G) \leq 1 - p_G \right) \\ &\Leftrightarrow \neg \exists \sigma' \in Adv_{M'}. \left(Pr_{M'}^{\sigma'}(\square \neg err_A)_{\geq p_A} \wedge Pr_{M'}^{\sigma'}(\diamond err_G) > 1 - p_G \right) \end{aligned}$$

where the second step follows from (1), the third since f is a bijection. \square

Proof .2 (of Theorem 3.5) The proof is by contradiction, therefore assume that there exist PAs M_1 and M_2 , regular safety properties A_1 , A_2 and G and probability bounds p_1 , p_2 and p_G such that $\langle A_1 \rangle_{\geq p_1} M_2 \langle G \rangle_{\geq p_G}$, $\langle A_2 \rangle_{\geq p_2} M_1 \langle A_1 \rangle_{\geq p_1}$ and $\langle true \rangle M_2 \langle A_2 \rangle_{\geq p_2}$ hold, while $\langle true \rangle M_1 \parallel M_2 \langle G \rangle_{\geq p_G}$ does not. From the latter, it follows that there exists an adversary $\sigma \in Adv_{M_1 \parallel M_2}$ such that $Pr_{M_1 \parallel M_2}^\sigma(G) < p_G$. Using the fact that both $\langle true \rangle M_2 \langle A_2 \rangle_{\geq p_2}$ and $\langle A_2 \rangle_{\geq p_2} M_1 \langle A_1 \rangle_{\geq p_1}$ hold, we can apply Theorem 3.2 to show that $\langle true \rangle M_1 \parallel M_2 \langle A_1 \rangle_{\geq p_1}$ holds, and hence:

$$\begin{aligned} Pr_{M_1 \parallel M_2}^\sigma(A_1) &\geq p_1 \\ \Rightarrow Pr_{M_2[\alpha_{A_1}]}^{\sigma \upharpoonright_{M_2[\alpha_{A_1}]}}(A_1) &\geq p_1 && \text{by Lemma 3.3(b) since } \alpha_{A_1} \subseteq \alpha_{M_2[\alpha_{A_1}]} \\ \Rightarrow Pr_{M_2[\alpha_{A_1}]}^{\sigma \upharpoonright_{M_2[\alpha_{A_1}]}}(G) &\geq p_G && \text{since } \langle A_1 \rangle_{\geq p_1} M_1 \langle G \rangle_{\geq p_G} \\ \Rightarrow Pr_{M_1 \parallel M_2}^\sigma(G) &\geq p_G && \text{by Lemma 3.3(b) since } \alpha_{A_1} \subseteq \alpha_{M_2[\alpha_{A_1}]} \end{aligned}$$

which contradicts the assumption that $Pr_{M_1 \parallel M_2}^\sigma(G) < p_G$. \square

Proof .3 (of Theorem 3.6) The proof is by contradiction. Suppose there exist PAs M_1 and M_2 where $\alpha_{M_1} \cap \alpha_{M_2} = \emptyset$, regular safety properties A_1 , A_2 , G_1 and G_2 and probability bounds p_{A_1} , p_{A_2} , p_{G_1} and p_{G_2} such that $\langle A_1 \rangle_{\geq p_{A_1}} M_1 \langle G_1 \rangle_{\geq p_{G_1}}$ and $\langle A_2 \rangle_{\geq p_{A_2}} M_2 \langle G_2 \rangle_{\geq p_{G_2}}$ hold while

$$\langle A_1, A_2 \rangle_{\geq p_{A_1}, p_{A_2}} M_1 \parallel M_2 \langle G_1 \vee G_2 \rangle_{\geq p_{G_1} + p_{G_2} - p_{G_1} \cdot p_{G_2}}$$

does not. Therefore, letting $\alpha_A = \alpha_{A_1} \cup \alpha_{A_2}$, by Definition 7 there exists an adversary σ of $(M_1 \parallel M_2)[\alpha_A]$ such that

$$Pr_{(M_1 \parallel M_2)[\alpha_A]}^\sigma(A_1) \geq p_{A_1} \quad (2)$$

$$Pr_{(M_1 \parallel M_2)[\alpha_A]}^\sigma(A_2) \geq p_{A_2} \quad (3)$$

$$Pr_{(M_1 \parallel M_2)[\alpha_A]}^\sigma(G_1 \vee G_2) < p_{G_1} + p_{G_2} - p_{G_1} \cdot p_{G_2}. \quad (4)$$

As G_1 and G_2 are safety properties, we have:

$$Pr_{(M_1 \parallel M_2)[\alpha_A]}^\sigma(G_1 \vee G_2) = 1 - Pr_{(M_1 \parallel M_2)[\alpha_A]}^\sigma \parallel (G_1^{err} \parallel G_2^{err}) (\diamond(err_{G_1} \wedge err_{G_2})). \quad (5)$$

Next, since $\alpha_A = \alpha_{A_1} \cup \alpha_{A_2}$ and the parallel composition operator is commutative and associative, using Definition 3 it follows that:

$$(M_1 \parallel M_2)[\alpha_A] \parallel (G_1^{err} \parallel G_2^{err}) = (M_1[\alpha_{A_1}] \parallel G_1^{err}) \parallel (M_2[\alpha_{A_2}] \parallel G_2^{err})$$

from which, using the fact that $\alpha_{M_1} \cap \alpha_{M_2} = \emptyset$, we can derive the equality:

$$\begin{aligned} & Pr_{(M_1 \parallel M_2)[\alpha_A]}^\sigma \parallel (G_1^{err} \parallel G_2^{err}) (\diamond(err_{G_1} \wedge err_{G_2})) \\ &= Pr_{M_1[\alpha_{A_1}] \parallel G_1^{err}}^{\sigma \upharpoonright_{M_1[\alpha_{A_1}]}} (\diamond err_{G_1}) \cdot Pr_{M_2[\alpha_{A_2}] \parallel G_2^{err}}^{\sigma \upharpoonright_{M_2[\alpha_{A_2}]}} (\diamond err_{G_2}). \end{aligned} \quad (6)$$

Since $\langle A_1 \rangle_{p_{A_1}} M_1 \langle G_1 \rangle_{\geq p_{G_1}}$ and $\langle A_2 \rangle_{p_{A_2}} M_2 \langle G_2 \rangle_{\geq p_{G_2}}$ hold, using (2) and (3) together with Definition 3 we have that:

$$Pr_{M_1[\alpha_{A_1}] \parallel G_1^{err}}^{\sigma \upharpoonright_{M_1[\alpha_{A_1}]}} (\diamond err_{G_1}) \leq 1 - p_{G_1} \quad \text{and} \quad Pr_{M_2[\alpha_{A_2}] \parallel G_2^{err}}^{\sigma \upharpoonright_{M_2[\alpha_{A_2}]}} (\diamond err_{G_2}) \leq 1 - p_{G_2}$$

which combined with (6) yields:

$$Pr_{(M_1 \parallel M_2)[\alpha_A]}^\sigma \parallel (G_1^{err} \parallel G_2^{err}) (\diamond(err_{G_1} \wedge err_{G_2})) \leq (1 - p_{G_1}) \cdot (1 - p_{G_2}).$$

Finally, substituting this result into (5) gives the inequality:

$$Pr_{(M_1 \parallel M_2)[\alpha_A]}^\sigma(G_1 \vee G_2) \geq 1 - (1 - p_{G_1}) \cdot (1 - p_{G_2}) = p_{G_1} + p_{G_2} - p_{G_1} \cdot p_{G_2}$$

which contradicts (4), and hence completes the proof. \square